

3.4. Maximum Likelihood Estimation

x_1, \dots, x_n independent sample from pdf
 $f(x, \vartheta)$, ϑ a parameter.

$$L(x; \vartheta) = \prod_{i=1}^n f(x_i, \vartheta) \quad \text{likelihood function}$$

$$l(x; \vartheta) = \log L(x; \vartheta) = \sum_{i=1}^n \log f(x_i, \vartheta)$$

log-likelihood fct.

Find ϑ which fits the data best, i.e.,

$$\hat{\vartheta} = \arg \max_{\vartheta} l(x; \vartheta)$$

$\hat{\vartheta}$ is called ML estimator. (MLE)

Th. 3.6. $X \sim N_p(\mu, \Sigma)$, x_1, \dots, x_n i.i.d. sample of X .

The MLEs of μ, Σ are:

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x}, \quad \hat{\Sigma} = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})(x_i - \bar{x})^T = S_n$$

Proof. Density of $N_p(\mu, \Sigma)$

$$f(x) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right), \quad x \in \mathbb{R}^p$$

$$l(x_1, \dots, x_n; \mu, \Sigma)$$

$$= \sum_{i=1}^n \left[\log \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} - \frac{1}{2} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) \right]$$

$$= \underbrace{n \log \frac{1}{(2\pi)^{p/2}} + \frac{n}{2} \log |\Sigma^{-1}|}_{\text{constant}} - \frac{1}{2} \sum_{i=1}^n (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

Leave the constant, set $\Lambda = \Sigma^{-1}$

$$\begin{aligned} l^*(\mu, \Sigma) &= \frac{n}{2} \log |\Lambda| - \frac{1}{2} \sum_{i=1}^n (x_i - \mu)^T \Lambda (x_i - \mu) \\ &= \frac{n}{2} \log |\Lambda| - \frac{1}{2} \sum_{i=1}^n \text{tr}(\Lambda (x_i - \mu)(x_i - \mu)^T) \\ &= \frac{n}{2} \log |\Lambda| - \frac{1}{2} \text{tr}\left(\Lambda \sum_{i=1}^n (x_i - \mu)(x_i - \mu)^T\right) \end{aligned}$$

Steiners rule:

$$\begin{aligned} &\sum_{i=1}^n (x_i - \mu)(x_i - \mu)^T \\ &= \underbrace{\sum_{i=1}^n (x_i - \bar{x})(x_i - \bar{x})^T}_{n S_n} + (\bar{x} - \mu)(\bar{x} - \mu)^T \\ &\geq n S_n \quad (\text{equality if } \mu = \bar{x}) \end{aligned}$$

$$\leq \frac{n}{2} \log |\Lambda| - \frac{n}{2} \text{tr}(\Lambda S_n) = l^*(\mu^*, \Lambda)$$

$$\max \ell^*(\mu^*, \Lambda)$$

$$\text{Need } \frac{\partial}{\partial \Lambda} \log |\Lambda| = (\Lambda^{-1})^T$$

$$\frac{\partial}{\partial \Lambda} \text{tr}(\Lambda A) = A^T$$

$$\frac{\partial}{\partial \Lambda} \ell^*(\mu^*, \Lambda) = \frac{1}{2} \Lambda^{-1} - \frac{1}{2} S_n \stackrel{!}{=} O_{p \times p}$$

$$\Leftrightarrow \Sigma^* = S_n \quad \square$$

4. Dimensionality Reduction

Represent data in a low dimensional space
high dim.

in an "optimal" way. Dim. 1, 2, 3 allow for
visualization.

4.1. Principal Component Analysis (PCA)

Loose as little information as possible.

Given data $x_1, \dots, x_n \in \mathbb{R}^p$.

a) Find a k -dim. subspace such that the
projections of x_1, \dots, x_n thereon represent the
data ~~at~~ on its best.

b) Preserve as much variance as possible.

a) and b) are equivalent. \rightarrow later

x_1, \dots, x_n independently sampled from some distribution.

Sample mean: $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$

Sample covariance matrix: $S_n = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})(x_i - \bar{x})^T$

$\left(\begin{array}{l} \bar{x} : \text{unbiased estimator of } E(X) \\ S_n : \text{unbiased estimator of } \text{Cov}(X) \end{array} \right)$

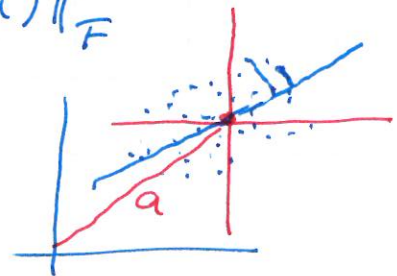
(Ex. MNIST data, $n=500$, $p=28 \cdot 28 = 784$)

4.1.1. Find the best projection.

Consider the opt. problem

$\min_{a \in \mathbb{R}^p} \sum_{i=1}^n \|x_i - a - Q(x_i - a)\|_F^2$

Q orth. proj. on a k -dim. subspace



$\min_{a, Q} \sum_{i=1}^n \|x_i - a - Q(x_i - a)\|^2$

$= \min_{a, Q} \sum_{i=1}^n \|(I - Q)(x_i - a)\|^2$

$= \min_{a, R} \sum_{i=1}^n \|R(x_i - a)\|^2, \quad R = I - Q \text{ (orth. proj. as well)}$

$$\begin{aligned}
 &= \min_{a, R} \sum_{i=1}^n (x_i - a)^T R^T R (x_i - a) \\
 &= \min_{a, R} \sum_{i=1}^n \text{tr}((x_i - a)^T R (x_i - a)) \\
 &= \min_{a, R} \sum_{i=1}^n \text{tr}(R (x_i - a)(x_i - a)^T) \\
 &= \min_{a, R} \text{tr}\left(R \sum_{i=1}^n (x_i - a)(x_i - a)^T\right) \\
 &\geq \min_R \text{tr}\left(R \sum_{i=1}^n (x_i - \bar{x})(x_i - \bar{x})^T\right) \quad \left(\begin{array}{l} \text{see MLE for } N(\mu, \Sigma) \\ \text{equality if } a = \bar{x} \end{array}\right) \\
 &= \min_R \text{tr}(R (n-1) S_n) \\
 &= \min_Q (n-1) \text{tr} S_n (I - Q)
 \end{aligned}$$

It remains to solve

$$\begin{aligned}
 \max_Q \text{tr}(S_n Q) \quad , \quad Q \text{ orth. proj.}, \quad Q = \sum_{i=1}^k q_i q_i^T, \quad q_i \text{ orth.} \\
 Q = \tilde{Q} \tilde{Q}^T, \quad \tilde{Q} = (q_1, \dots, q_k) \\
 = \max_{\tilde{Q}^T \tilde{Q} = I_k} \text{tr}(\tilde{Q}^T S_n \tilde{Q}) = \sum_{i=1}^k \lambda_i(S_n) \quad (\text{Ky Fan, Th. 2.4})
 \end{aligned}$$

where $\lambda_1(S_n) \geq \dots \geq \lambda_k(S_n) \geq \dots \geq \lambda_p(S_n)$ are the eigenvalues of S_n in decreasing order.

The max is attained if q_1, \dots, q_k are the orthogonal eigenvector corresponding to $\lambda_1(S_n), \dots, \lambda_k(S_n)$.

4.1.2 Preserve most variance

Seek a hyperplane so that the proj. data has most variance.

$$\max_Q \sum_{i=1}^n \|Qx_i - \frac{1}{n} \sum_{l=1}^n Qx_l\|^2$$

$$\max_Q \sum_{i=1}^n \|Qx_i - \frac{1}{n} \sum_{l=1}^n Qx_l\|^2, \quad Q = \tilde{Q}\tilde{Q}^T, \\ \tilde{Q}^T\tilde{Q} = I_R \\ \text{Orth. proj.}$$

$$= \max_Q \sum_{i=1}^n \|Qx_i - Q\bar{x}\|^2$$

$$= \max_Q \sum_{i=1}^n \|Q(x_i - \bar{x})\|^2$$

$$= \max_Q \sum_{i=1}^n \text{tr}(x_i - \bar{x})^T Q (x_i - \bar{x})$$

$$= \max_Q \text{tr} Q \sum_{i=1}^n (x_i - \bar{x})(x_i - \bar{x})^T$$

$$= \max_Q (n-1) \text{tr} Q S_n$$

with the same solution as above.

4.1.3 How to carry out PCA

Given $x_1, \dots, x_n \in \mathbb{R}^p$, fix $k \ll p$

Compute
$$S_n = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})(x_i - \bar{x})^T$$

$$S_n = V \Lambda V^T, \quad \Lambda = \text{diag}(\lambda_1, \dots, \lambda_p)$$

$\lambda_1 \geq \dots \geq \lambda_p$, $V = (v_1, \dots, v_p) \in \mathcal{O}(p)$ spectral decomposition

v_1, \dots, v_k are called the k principal eigenvectors to the principal eigenvalues $\lambda_1, \dots, \lambda_k$.

Projected points $\hat{x}_i = \begin{pmatrix} v_1^T \\ \vdots \\ v_k^T \end{pmatrix} x_i, \quad i=1, \dots, n.$

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